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# Reductions in hospital care among clinically vulnerable children aged 0–4 years during the COVID-19 pandemic

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## ABSTRACT

**Objective** To quantify reductions in hospital care for clinically vulnerable children during the COVID-19 pandemic.

**Design** Birth cohort.

**Setting** National Health Service hospitals in England.

**Study population** All children aged <5 years with a birth recorded in hospital administrative data (January 2010–March 2021).

**Main exposure** Clinical vulnerability defined by a chronic health condition, preterm birth (<37 weeks' gestation) or low birth weight (<2500 g).

**Main outcomes** Reductions in care defined by predicted hospital contact rates for 2020, estimated from 2015 to 2019, minus observed rates per 1000 child years during the first year of the pandemic (March 2020–2021).

**Results** Of 3 813 465 children, 17.7% (one in six) were clinically vulnerable (9.5% born preterm or low birth weight, 10.3% had a chronic condition). Reductions in hospital care during the pandemic were much higher for clinically vulnerable children than peers: respectively, outpatient attendances (314 vs 73 per 1000 child years), planned admissions (55 vs 10) and unplanned admissions (105 vs 79). Clinically vulnerable children accounted for 50.1% of the reduction in outpatient attendances, 55.0% in planned admissions and 32.8% in unplanned hospital admissions. During the pandemic, weekly rates of planned care returned to prepandemic levels for infants with chronic conditions but not older children. Reductions in care differed by ethnic group and level of deprivation. Virtual outpatient attendances increased from 3.2% to 24.8% during the pandemic.

**Conclusion** One in six clinically vulnerable children accounted for one-third to one half of the reduction in hospital care during the pandemic.

## INTRODUCTION

Rates of hospital contact (outpatient attendances, planned and unplanned hospital admissions) among 0–4 year-olds are highest for infants and have increased steadily in England over the past decade.<sup>1–3</sup> Hospital utilisation patterns differ markedly by age and clinical vulnerability: children born preterm, with low birth weight or a congenital anomaly, have substantially more admissions than other children.<sup>4 5</sup>

Hospital contacts reduced substantially during the COVID-19 lockdown,<sup>6–9</sup> which likely most impacted high intensity users of planned hospital

## WHAT IS ALREADY KNOWN ON THIS TOPIC

- ⇒ Young children with chronic health conditions have high rates of hospital care.
- ⇒ Rates of planned and unplanned hospital admissions and outpatient care rose steadily in the decade before the pandemic.
- ⇒ Hospital care reduced sharply after the onset of the COVID-19 pandemic for all children.

## WHAT THIS STUDY ADDS

- ⇒ One in six clinically vulnerable children 0–4 years old accounted for one-third to one-half of reductions in hospital care during the pandemic in this age group.
- ⇒ There were small differences in reductions between black and Asian ethnic groups (vs white) and for children in the most deprived quintile (vs least deprived).
- ⇒ Planned care weekly rates (outpatient or admissions) for children with chronic conditions reduced sharply during the pandemic and returned to prepandemic levels only among infants.

## HOW THIS STUDY MIGHT AFFECT RESEARCH, PRACTICE OR POLICY

- ⇒ Research is needed to understand the reductions in planned care, the types of care or procedures affected and the short-term and long-term implications for children.
- ⇒ Research is needed to ascertain whether contact rates have returned to previous levels since the ending of pandemic restrictions.
- ⇒ Children in vulnerable groups will likely need targeted catch-up funding and resources to mitigate or prevent adverse outcomes accruing from the reductions we report.

care, such as children born too early or too small or with underlying health conditions.<sup>1 2 10</sup> Postponed or cancelled planned hospital care may result in delayed diagnoses or treatments, which could be detrimental to health or development.<sup>11</sup> Fewer unplanned hospital admissions might reflect fewer infections, injuries or other health problems due to reduced exposure during pandemic restrictions but could also reflect unmet need.

This study aimed to quantify reductions in planned and unplanned hospital care for clinically vulnerable children and non-vulnerable peers



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during the COVID-19 pandemic using national, longitudinal administrative hospital data for England. We measured planned hospital contacts (admissions and outpatient attendances) and unplanned admissions among children with chronic health conditions or born preterm or low birth weight and children with no recorded clinical vulnerability as well as by ethnicity and area-based deprivation. Reductions were quantified as the difference between predicted and observed rates of hospital contacts during the first year of the pandemic. We also examined whether rates of contact returned to prepandemic levels and described changes in the type of outpatient contact (eg, face to face or virtual).

## METHODS

### Study population and data source

Children were included in the cohort if aged 0–4 years and their birth was recorded in hospital administrative data in the English NHS (Hospital Episode Statistics (HES)) between 1 January 2010 and 31 March 2021. HES records 97% of all births in England.<sup>12</sup> Children were followed until the earliest of fifth birthday or 31 March 2021 (online supplemental figure 1). All contacts with NHS hospitals in England (including admitted patient care and outpatient attendances) were linked using the encrypted HES Identifier (ID). Accident and emergency (A&E) attendances could not be included as patient-level linkage is via Token ID.<sup>13</sup> We combined consecutive consultant episodes and hospital transfers to form admissions.<sup>14</sup>

### Outcome and exposure

Our primary outcome was the reduction in hospital care, defined as the difference between observed and predicted rates (per 1000 child-years (cy)) of hospital contacts (stratified into outpatient attendances, planned or unplanned admissions) during the pandemic (23 March 2020–22 March 2021). Inpatient admissions were classified as planned or unplanned using the ‘admission method’ variable (admissions have a predictable clinical need where the decision to admit may be separated in time from the actual admission).<sup>15</sup> We also described trends in weekly rates of hospital contacts by age (from 1 January 2020 to 31 March 2021 and averaged for 2015–2019), and uptake (attended, missed, cancelled and postponed) and mode (in person, virtually) of outpatient contacts.

Clinical vulnerability was defined by a chronic condition recorded up to age 4 years, or preterm birth (<37 weeks of gestation) or low birth weight (<2500 g) recorded in birth or delivery records. A child had a chronic condition at a given age if at least one relevant International Classification of Diseases 10th revision (ICD-10) code (identified using code lists developed by Feudtner and Hardelid<sup>12 16–18</sup>) was recorded in any of their records up to that age including their birth admission. Completeness of demographic variables declined for children born during the pandemic (missingness increase by; 9% Index of Multiple Deprivation (IMD); 3% birth weight; 2% gestational age) but was unchanged for admission characteristics including inpatient admission type (planned, unplanned, birth or maternity) or outpatient attendance type (in-person and virtual), similar to outpatient data metrics published by NHS digital.<sup>19</sup> Children with a chronic condition but missing gestation and birth weight were categorised as clinically vulnerable (chronic condition only).

We analysed risk factors associated with health service use: age (0–11 months, 1–4 years),<sup>3</sup> quintile/fifth of deprivation derived from the national distribution of the IMD 2004 (an area measure

for ~650 households)<sup>20 21</sup> and recorded ethnic group (grouped as white, black, Asian or other, including mixed and Chinese).

### Statistical analyses

We calculated observed rates of hospital contacts per 1000 cy in the prepandemic period (2015–2019), stratifying by risk factors.

We calculated child-years at risk by averaging the eligible population of births recorded in HES at the beginning and end of a year, assuming no emigration and ignoring deaths. We used Poisson regression (including a linear effect of time and log of the midyear population as an offset) to model rates stratified by risk factors from 1 January 2015 to 31 December 2019. Data from 1 January to 22 March 2020 were excluded from the prepandemic period because reductions in hospital contact rates preceded the first lockdown. To calculate the reduction in hospital contacts, we predicted rates for the pandemic period, assuming that the pandemic had not occurred and previous trends would have continued. The reduction was estimated as the difference between predicted and observed rates. We also calculated reductions within the first national lockdown (23 March–23 June 2020), easing of restrictions (24 June–4 November 2020), second national lockdown (5 November–31 December 2020) and third national lockdown (1 January–22 March 2021).<sup>22</sup>

We calculated weekly rates between 1 January 2015 and 31 March 2021, using dynamic denominator study populations to accommodate temporal changes in hospital activity. Analyses stratified children with and without chronic conditions (recorded between birth and the relevant week) because we expected the differences to be largest between these comparators. Weekly rates of hospital contact were calculated by dividing the total number of weekly admissions or attendances by the weekly dynamic denominator population of children within each stratification level (ie, a child born in week 1 of 2015 would move into the 1-year-old group in week 1 of 2016 and age out of the cohort in week 1 of 2019). Weekly rates in 2020 and 2021 were plotted against average weekly rates for 2015–2019. We also modelled weekly rates between 1 January 2015 and 31 December 2019, using a Poisson model that included a linear effect of time, calendar month to account for seasonality, log of the weekly denominator population as an offset and second-order lagged residuals. A similar approach was used to estimate weekly reductions during the pandemic. Analyses were performed in Stata V.16.<sup>23</sup>

## RESULTS

### Population characteristics

Of the 3 813 465 children aged 0–4 years, 394 384 (10.3%) had a record indicating a chronic condition (including congenital anomalies); 363 950 (9.5%) were born preterm or low birth weight; and 83 283 (2.2%) had both vulnerabilities. Overall, 675 051 (17.7%) had one or more of these clinical vulnerabilities (table 1).

### Hospital contacts prepandemic

Hospital contacts were much higher among infants than children aged 1–4 years: 60.1% (95% CI 60.0 to 60.3) of infants and 8.2% (8.1–8.2) of 1–4 year-olds had ≥1 outpatient attendance each year, reflecting 1538 outpatient attendances per 1000 cy for infants and 302 for 1–4 year-olds (online supplemental figure 2, online supplemental table 1). Overall, 31.2% (31.1–31.3) of clinically vulnerable children had ≥1 outpatient attendance compared with 14.1% (14.1–14.2) of those with no known

**Table 1** Demographic characteristics of children born between 1 January 2015 and 31 March 2021 by vulnerability status

|                         | None<br>(n=2 773 420; 72.7%) |      | LBW/preterm only<br>(n=280 667; 7.4%) |      | CC only<br>(n=311 101; 8.25) |      | Both<br>(n=83 283; 2.2%) |      | Missing<br>(n=364 994; 9.6%) |      | Total<br>(n=3 813 465) |      |
|-------------------------|------------------------------|------|---------------------------------------|------|------------------------------|------|--------------------------|------|------------------------------|------|------------------------|------|
|                         | n                            | %    | n                                     | %    | n                            | %    | n                        | %    | n                            | %    | n                      | %    |
| <b>Sex</b>              |                              |      |                                       |      |                              |      |                          |      |                              |      |                        |      |
| Male                    | 1 406 322                    | 50.7 | 134 871                               | 48.1 | 183 964                      | 59.1 | 46 304                   | 55.6 | 183 986                      | 50.4 | 1 955 447              | 51.3 |
| Female                  | 1 366 276                    | 49.3 | 145 481                               | 51.8 | 127 006                      | 40.8 | 36 860                   | 44.3 | 180 806                      | 49.5 | 1 856 429              | 48.7 |
| Missing                 | 822                          | 0.03 | 315                                   | 0.1  | 131                          | 0.04 | 119                      | 0.1  | 202                          | 0.1  | 1589                   | 0.04 |
| <b>Age</b>              |                              |      |                                       |      |                              |      |                          |      |                              |      |                        |      |
| Infants                 | 400 330                      | 14.4 | 36 445                                | 13.0 | 27 641                       | 8.9  | 8317                     | 10.0 | 80 502                       | 22.1 | 553 235                | 14.5 |
| 1–4 years old           | 2 373 090                    | 85.6 | 244 222                               | 87.0 | 283 460                      | 91.1 | 74 966                   | 90.0 | 284 492                      | 77.9 | 3 260 230              | 85.5 |
| <b>Ethnicity</b>        |                              |      |                                       |      |                              |      |                          |      |                              |      |                        |      |
| White                   | 1 958 159                    | 70.6 | 184 525                               | 65.8 | 228 023                      | 73.3 | 58 366                   | 70.1 | 227 775                      | 62.4 | 2 656 848              | 69.7 |
| Black/black British     | 126 067                      | 4.5  | 15 477                                | 5.5  | 14 150                       | 4.6  | 4681                     | 5.6  | 14 924                       | 4.1  | 175 299                | 4.6  |
| Asian/Asian British     | 295 850                      | 10.7 | 41 516                                | 14.8 | 33 610                       | 10.8 | 10 857                   | 13.0 | 36 159                       | 9.9  | 417 992                | 11.0 |
| Any other ethnic groups | 239 079                      | 8.6  | 24 085                                | 8.6  | 25 135                       | 8.1  | 6633                     | 8.0  | 34 579                       | 9.5  | 329 511                | 8.6  |
| Missing                 | 154 265                      | 5.6  | 15 064                                | 5.4  | 10 183                       | 3.3  | 2746                     | 3.3  | 51 557                       | 14.1 | 233 815                | 6.1  |
| <b>IMD</b>              |                              |      |                                       |      |                              |      |                          |      |                              |      |                        |      |
| Q1 (most deprived)      | 668 113                      | 24.1 | 83 204                                | 29.7 | 89 425                       | 28.7 | 27 254                   | 32.7 | 71 436                       | 19.6 | 939 432                | 24.6 |
| Q2                      | 550 124                      | 19.8 | 60 337                                | 21.5 | 67 853                       | 21.8 | 18 629                   | 22.4 | 62 541                       | 17.1 | 759 484                | 19.9 |
| Q3                      | 468 995                      | 16.9 | 45 794                                | 16.3 | 55 924                       | 18.0 | 14 117                   | 17.0 | 55 868                       | 15.3 | 640 698                | 16.8 |
| Q4                      | 413 514                      | 14.9 | 37 446                                | 13.3 | 47 670                       | 15.3 | 11 287                   | 13.6 | 47 604                       | 13.0 | 557 521                | 14.6 |
| Q5 (least deprived)     | 400 528                      | 14.4 | 34 530                                | 12.3 | 43 052                       | 13.8 | 10 119                   | 12.2 | 35 619                       | 9.8  | 523 848                | 13.7 |
| Missing                 | 272 146                      | 9.8  | 19 356                                | 6.9  | 7177                         | 2.3  | 1877                     | 2.3  | 91 926                       | 25.2 | 392 482                | 10.3 |
| <b>Birth year</b>       |                              |      |                                       |      |                              |      |                          |      |                              |      |                        |      |
| 2015                    | 457 046                      | 16.5 | 46 590                                | 16.6 | 68 814                       | 22.1 | 16 021                   | 19.2 | 55 578                       | 15.2 | 644 049                | 16.9 |
| 2016                    | 471 324                      | 17.0 | 49 020                                | 17.5 | 63 505                       | 20.4 | 16 042                   | 19.3 | 44 186                       | 12.1 | 644 077                | 16.9 |
| 2017                    | 465 975                      | 16.8 | 49 358                                | 17.6 | 55 396                       | 17.8 | 15 096                   | 18.1 | 45 863                       | 12.6 | 631 688                | 16.6 |
| 2018                    | 443 438                      | 16.0 | 47 406                                | 16.9 | 47 628                       | 15.3 | 13 730                   | 16.5 | 56 695                       | 15.5 | 608 897                | 16.0 |
| 2019                    | 431 854                      | 15.6 | 41 713                                | 14.9 | 40 107                       | 12.9 | 11 646                   | 14.0 | 65 828                       | 18.0 | 591 148                | 15.5 |
| 2020                    | 409 956                      | 14.8 | 38 128                                | 13.6 | 30 598                       | 9.8  | 9341                     | 11.2 | 77 409                       | 21.2 | 565 432                | 14.8 |
| 2021                    | 93 827                       | 3.4  | 8452                                  | 3.0  | 5053                         | 1.6  | 1407                     | 1.7  | 19 435                       | 5.3  | 128 174                | 3.4  |

CC, chronic conditions; IMD, Index of Multiple Deprivation; LBW, low birth weight.

vulnerability reflecting 1483 attendances per 1000 cy for children with any vulnerability and 295/1000 cy for those without. Patterns were similar for planned and unplanned hospital admissions. Children with chronic conditions had the highest rates of admissions across all strata (online supplemental figures 3 and 4). Children born preterm or low birth weight but with no chronic condition had similar admission rates to their peers born at term or weighing  $\geq 2500$  g (online supplemental figure 2).

### Hospital contacts during the pandemic

There were stark reductions in rates of all hospital contact types during the pandemic (table 2, online supplemental tables 2–4). Reductions were much larger for children with a chronic condition (outpatient:  $-492$  ( $-505$  to  $-480$ ) contacts per 1000cy; planned:  $-91/1000$  cy ( $-95$  to  $-86$ ); unplanned:  $-230/1000$  cy ( $-236$  to  $-224$ )) than those without and particularly high for children with a chronic condition who were also born preterm or with a low birth weight (outpatient:  $-536/1000$  cy ( $-563$  to  $-508$ ); planned:  $-105/1000$  cy ( $-113$  to  $-97$ ); unplanned:  $-279/1000$  cy ( $-290$  to  $-268$ )). Relative reductions were lower for children with any known vulnerability compared with those without any known vulnerability (outpatient:  $-19.0\%$  ( $-19.0$  to  $-18.9$ ) vs  $-25.7\%$  ( $-25.8$  to  $-25.6$ ); planned:  $-27.6\%$  ( $-27.9$  to  $-27.3$ ) vs  $-58.0\%$  ( $-58.5$  to  $-57.6$ ); unplanned:  $-45.9\%$  ( $-46.1$  to  $-45.7$ ) vs  $-69.4\%$  ( $-69.6$  to  $-69.3$ )) (table 2). The 17.7% of children identified as clinically vulnerable accounted for 50.1% (49.9–50.3) of reductions in outpatient attendances,

55.0% (54.6–55.5) in planned hospital admissions and 32.8% (32.6–33.0) in unplanned hospital admissions (table 2).

Reductions were larger for infants than 1–4 year-olds for outpatient attendances and unplanned admissions but not for planned admissions (figure 1). We found small differences in reductions of planned and unplanned admissions across ethnic groups and in all hospital contacts among children in the most (vs least) deprived quintile (online supplemental tables 3 and 4). The largest reductions in care were among children with a chronic condition (online supplemental figures 5 and 6). Overall, the first lockdown was associated with the largest reductions in outpatient attendances and planned admissions. The largest reductions in unplanned admissions were seen in the second lockdown (table 3; online supplemental tables 5–7).

### Trends in hospital contacts

Outpatient attendances reduced sharply before and during the first national lockdown, among children of all ages with a chronic condition, with less perceptible changes among those without a chronic condition. Outpatient attendances rapidly returned to prepandemic rates for infants but remained below 2015–19 averages for older children. Planned admissions followed a similar pattern, with a return to prepandemic rates only for infants (figure 2).

A similar pattern was observed for reductions in rates of unplanned admissions, and these remained below prepandemic levels for both groups at all ages (figure 2). In 2020, the

**Table 2** Difference between predicted and observed rates of hospital contact during the pandemic (March 2020–2021) among children aged 0–4 years by clinical vulnerability group

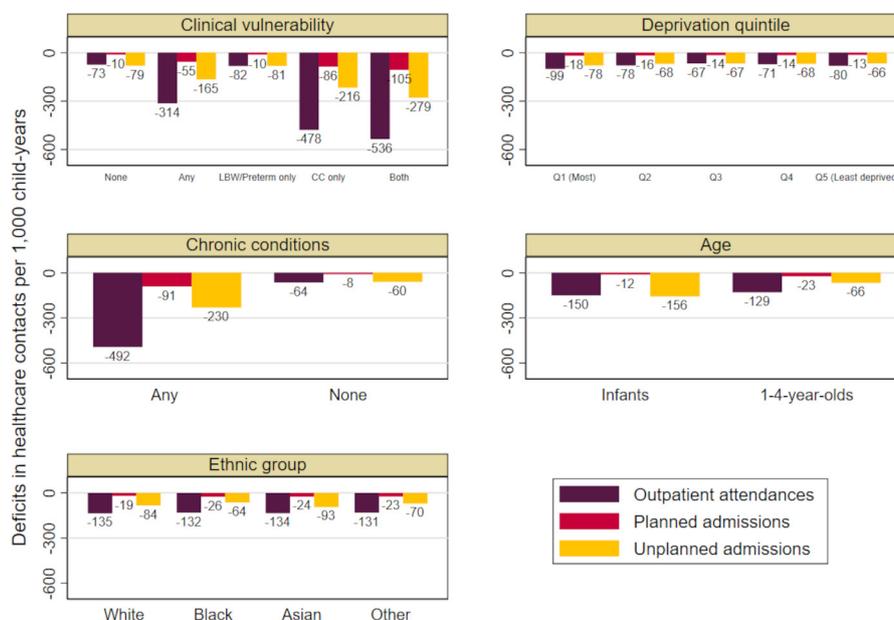
|                               | Percentage of children seen |           |      | Number of hospital contacts |           |            |          | Rates per 1000 child-years |          |                     |          |
|-------------------------------|-----------------------------|-----------|------|-----------------------------|-----------|------------|----------|----------------------------|----------|---------------------|----------|
|                               | n                           | N         | %    | Predicted                   | Observed  | Difference | % change | Predicted                  | Observed | Difference (95% CI) | % change |
| <b>Outpatient attendances</b> |                             |           |      |                             |           |            |          |                            |          |                     |          |
| Total*                        | 370623                      | 2 765 941 | 13.4 | 1 500 674                   | 1 173 091 | -327 583   | -21.8    | 543                        | 424      | -118 (-115 to 122)  | -21.8    |
| No known vulnerability        | 232 976                     | 2 243 088 | 10.4 | 635 226                     | 471 743   | -163 483   | -25.7    | 283                        | 210      | -73 (-71 to 74)     | -25.7    |
| Any vulnerability             | 137 647                     | 522 853   | 26.3 | 865 448                     | 701 348   | -164 100   | -19.0    | 1655                       | 1341     | -314† (-302 to 326) | -19.0    |
| LBW/preterm only              | 29 622                      | 226 439   | 13.1 | 94 660                      | 76 039    | -18 621    | -19.7    | 418                        | 336      | -82† (-77 to 88)    | -19.7    |
| CC only                       | 86 825                      | 231 671   | 37.5 | 586 026                     | 475 226   | -110 800   | -18.9    | 2530                       | 2051     | -478† (-465 to 492) | -18.9    |
| Both                          | 21 200                      | 64 743    | 32.7 | 184 762                     | 150 083   | -34 679    | -18.8    | 2854                       | 2318     | -536† (-508 to 563) | -18.8    |
| Any CC‡                       | 107 762                     | 296 414   | 36.4 | 771 219                     | 625 309   | -145 910   | -18.9    | 2602                       | 2110     | -492† (-480 to 505) | -18.9    |
| <b>Planned admissions</b>     |                             |           |      |                             |           |            |          |                            |          |                     |          |
| Total                         | 52 280                      | 2 765 941 | 1.9  | 144 980                     | 92 618    | -52 362    | -36.1    | 52                         | 33       | -19 (-18 to 20)     | -36.1    |
| No known vulnerability        | 15 287                      | 2 243 088 | 0.7  | 40 568                      | 17 022    | -23 546    | -58.0    | 18                         | 8        | -10 (-10 to 11)     | -58.0    |
| Any vulnerability             | 36 993                      | 522 853   | 7.1  | 104 412                     | 75 596    | -28 816    | -27.6    | 200                        | 145      | -55† (-51 to 59)    | -27.6    |
| LBW/preterm only              | 1389                        | 226 439   | 0.6  | 3675                        | 1522      | -2153      | -58.6    | 16                         | 7        | -10 (-8 to 11)      | -58.6    |
| CC only                       | 31 038                      | 231 671   | 13.4 | 85 367                      | 65 506    | -19 861    | -23.3    | 368                        | 283      | -86† (-81 to 91)    | -23.3    |
| Both                          | 4566                        | 64 743    | 7.1  | 15 370                      | 8568      | -6802      | -44.3    | 237                        | 132      | -105† (-97 to 113)  | -44.3    |
| Any CC                        | 35 603                      | 296 414   | 12.0 | 100 929                     | 74 074    | -26 855    | -26.6    | 341                        | 250      | -91† (-86 to 95)    | -26.6    |
| <b>Unplanned admissions</b>   |                             |           |      |                             |           |            |          |                            |          |                     |          |
| Total                         | 131 134                     | 2 765 941 | 4.7  | 442 174                     | 179 366   | -262 808   | -59.4    | 160                        | 65       | -95 (-93 to 97)     | -59.4    |
| No known vulnerability        | 66 830                      | 2 243 088 | 3.0  | 254 263                     | 77 713    | -176 550   | -69.4    | 113                        | 35       | -79 (-78 to 80)     | -69.4    |
| Any vulnerability             | 64 304                      | 522 853   | 12.3 | 187 911                     | 101 653   | -86 258    | -45.9    | 359                        | 194      | -165† (-159 to 171) | -45.9    |
| LBW/preterm only              | 6046                        | 226 439   | 2.7  | 25 520                      | 7228      | -18 292    | -71.7    | 113                        | 32       | -81 (-78 to 84)     | -71.7    |
| CC only                       | 52 285                      | 231 671   | 22.6 | 134 055                     | 84 126    | -49 929    | -37.2    | 579                        | 363      | -216† (-209 to 222) | -37.2    |
| Both                          | 5973                        | 64 743    | 9.2  | 28 336                      | 10 299    | -18 037    | -63.7    | 438                        | 159      | -279† (-268 to 290) | -63.7    |
| Any CC                        | 58 248                      | 296 414   | 19.7 | 162 648                     | 94 425    | -68 223    | -41.9    | 549                        | 319      | -230† (-224 to 236) | -41.9    |

\*268 256 (8.8%) children missing gestational age and birth weight data.  
 †Significantly different from children with no known vulnerability (5% level of significance).  
 ‡Any CC combines CC only and both.  
 CC, chronic conditions; LBW, low birth weight.

autumn–winter peak in unplanned admissions was diminished relative to previous years; however, following the reopening of primary schools at the end of the third lockdown on 8 March 2021, there was an increase in unplanned admission rates for all children, particularly those with a chronic condition (figure 2,

online supplemental figure 7). Trends did not consistently differ across deprivation levels (online supplemental figures 8 and 9) or by ethnic group (online supplemental figures 10–12).

For all age groups, a spike in cancellations and postponement of outpatient appointments preceded the first lockdown by



**Figure 1** Reduction in care during the pandemic (March 2020–2021), estimated from predicted minus observed rates of hospital contacts per 1000 child-years for children aged 0–4 years, by clinical vulnerability status and risk factors. CC, chronic conditions; LBW, low birth weight.

3 weeks (online supplemental figure 13). There was an increase in tele/virtual outpatient attendances during the pandemic and face-to-face visits did not return to prepandemic levels in any age group (online supplemental table 8 and figure 14).

## DISCUSSION

This population-based cohort study of all children aged <5 years in England found large and disproportionate reductions in planned and unplanned hospital contacts during the COVID-19 pandemic for clinically vulnerable groups. The one in six clinically vulnerable children accounted for over half the reduction in outpatient attendances and planned admissions, and one-third of the reduction in unplanned admissions. While absolute reductions were larger for vulnerable children, they were smaller in relative terms compared with children with no recorded vulnerability suggesting that hospitals prioritised these children. We saw some evidence of recovery in planned care during the pandemic among infants, but not among older children.

This study's main strength is the use of a birth cohort of all children born in an NHS hospital in England (97% of all births). This large sample size gave us enough data to calculate weekly rates of hospital contacts. The longitudinal nature of the data allowed us to identify chronic conditions from diagnostic codes recorded in all admissions since birth, using a clinically developed coding system.<sup>16</sup>

Limitations include underascertainment of chronic conditions for children who could not be admitted to hospital due to the pandemic. These children may have been managed in primary care, or as outpatients, where chronic conditions coding is mostly missing. Furthermore, older children would have had more time

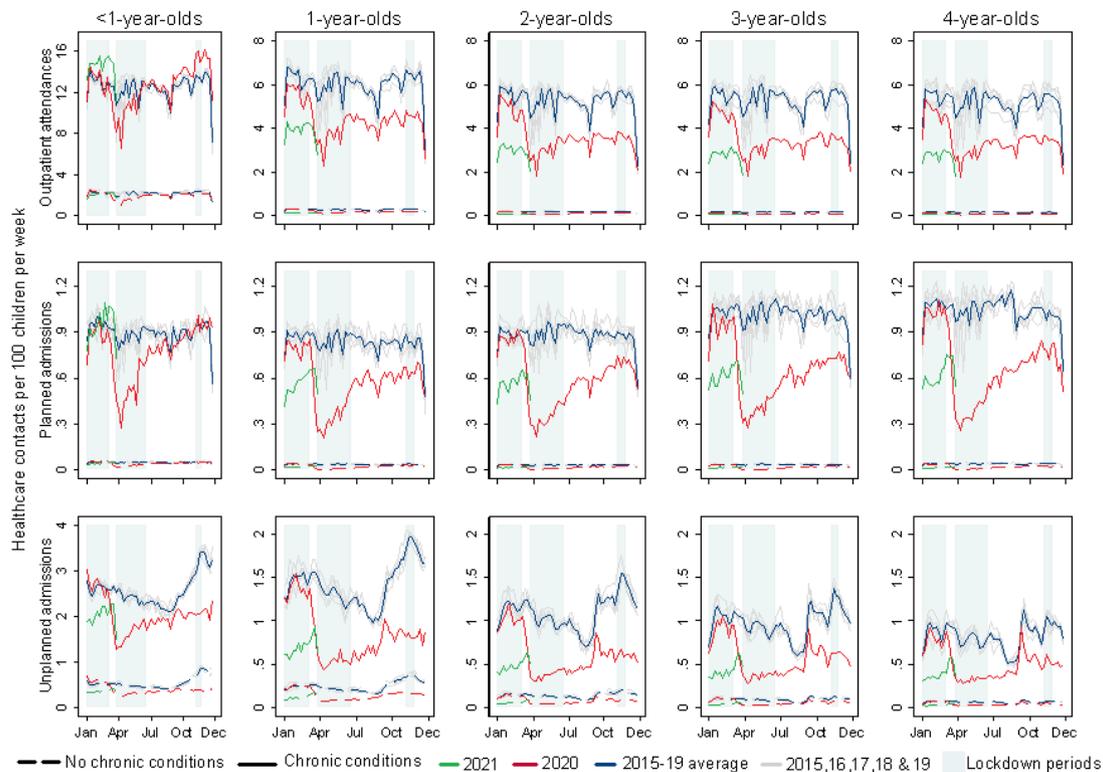
for chronic conditions to be diagnosed. This likely explains the decline in prevalence of chronic conditions in more recent years. Analyses were restricted by variables available in HES (eg, use of IMD 2004). Vulnerability may be underestimated for the 10% of children who were excluded from the study due to missing gestational age and birth weight data. Multiple imputation of missing data was not feasible given the study size. We could not quantify the reduction in A&E attendances, as this dataset is not yet linkable to admission records. However, recent studies investigating A&E attendances in children reported similar reduced service use during the pandemic.<sup>24 25</sup> Our rates did not account for deaths in the denominator (0.5%), non-NHS healthcare or emigration, but these events are rare. Our modelling approach required several assumptions, and our estimates of impacts are likely conservative.

Reductions in hospital care for children during the pandemic have been reported in Europe,<sup>7 26–33</sup> Asia,<sup>34</sup> North<sup>35–38</sup> and South America.<sup>39</sup> Most studies investigated A&E attendances or unplanned admissions.<sup>25–39</sup> Other studies report a reduction in asthma-related paediatric emergency department attendances<sup>29</sup> and reduced likelihood of admission, assessment and surgery for children with epilepsy.<sup>40</sup> Furthermore, significant reductions in infection-related hospitalisations have been observed,<sup>30 34 36 41</sup> particularly for children under 5 years.<sup>41</sup> Two studies conducted national level analyses.<sup>30 32</sup> We believe our study reports the first population-level reductions in planned care (admissions and outpatient) for children with health risk-factors for a full year of the pandemic. Previous research in adult populations has reported on the disproportionate burden of COVID-19 infection, hospitalisation and death in minority ethnic groups.<sup>42 43</sup>

**Table 3** Difference in predicted and observed rates of hospital contact per 1000 child-years among children aged 0–4 years during the pandemic (March 2020–2021) by period and clinical vulnerability

|                        | First lockdown<br>23 March–23 June 2020 |          | Easing of restrictions<br>24 June–4 November 2020 |          | second lockdown<br>Nov five to Dec 31, 2020 |          | third lockdown<br>Jan one to March 22, 2021 |          |
|------------------------|---|----------|---|----------|---|----------|---|----------|
|                        | Reduction (95% CI)                      | % change | Reduction (95% CI)                                | % change | Reduction (95% CI)                          | % change | Reduction (95% CI)                          | % change |
| Outpatient attendances |   |          |   |          |   |          |   |          |
| Total                  | -167 (-163 to 170)                      | -30.7    | -88 (-86 to 91)                                   | -16.3    | -96 (-91 to 100)                            | -17.9    | -128 (-123 to 132)                          | -23.5    |
| No known vulnerability | -106 (-104 to 107)                      | -37.2    | -55 (-54 to 56)                                   | -19.5    | -56 (-55 to 58)                             | -20.6    | -75 (-73 to 77)                             | -26.2    |
| Any vulnerability      | -428* (-417 to 439)                     | -25.9    | -230* (-221 to 240)                               | -13.9    | -264* (-249 to 278)                         | -16.0    | -352* (-337 to 367)                         | -21.4    |
| LBW/preterm only       | -121* (-116 to 127)                     | -28.7    | -57 (-53 to 61)                                   | -13.8    | -73* (-67 to 80)                            | -18.0    | -84 (-77 to 92)                             | -19.9    |
| CC only                | -664* (-651 to 677)                     | -26.2    | -357* (-346 to 368)                               | -14.0    | -398* (-381 to 414)                         | -15.9    | -516* (-499 to 533)                         | -20.8    |
| Both                   | -655* (-629 to 681)                     | -23.3    | -385* (-364 to 407)                               | -13.7    | -450* (-416 to 483)                         | -15.5    | -699* (-665 to 735)                         | -24.2    |
| Any CC                 | -663* (-652 to 675)                     | -25.5    | -364* (-354 to 374)                               | -14.0    | -410* (-395 to 425)                         | -15.8    | -559* (-544 to 574)                         | -21.7    |
| Planned admissions     |   |          |   |          |   |          |   |          |
| Total                  | -32 (-31 to 33)                         | -59.2    | -16 (-15 to 17)                                   | -29.3    | -9 (-8 to 11)                               | -19.2    | -16 (-15 to 17)                             | -31.9    |
| No known vulnerability | -14 (-14 to 15)                         | -79.1    | -9 (-9 to 9)                                      | -48.5    | -7 (-6 to 7)                                | -38.9    | -11 (-11 to 12)                             | -62.0    |
| Any vulnerability      | -105* (-101 to 109)                     | -51.6    | -46* (-43 to 49)                                  | -22.0    | -22* (-18 to 26)                            | -11.6    | -36* (-32 to 41)                            | -19.4    |
| LBW/preterm only       | -12* (-11 to 13)                        | -75.3    | -8 (-7 to 9)                                      | -49.4    | -6 (-5 to 7)                                | -42.1    | -11 (-11 to 10)                             | -64.9    |
| CC only                | -185* (-180 to 190)                     | -49.0    | -71* (-67 to 75)                                  | -18.4    | -24* (-18 to 30)                            | -6.9     | -40* (-34 to 46)                            | -11.6    |
| Both                   | -144* (-137 to 152)                     | -60.3    | -86* (-80 to 93)                                  | -35.8    | -70* (-61 to 79)                            | -31.6    | -114* (-105 to 125)                         | -48.2    |
| Any CC                 | -177* (-172 to 181)                     | -50.8    | -75* (-71 to 78)                                  | -21.2    | -35* (-30 to 40)                            | -10.7    | -57* (-52 to 62)                            | -17.7    |
| Unplanned admissions   |   |          |   |          |   |          |   |          |
| Total                  | -88 (-87 to 90)                         | -60.0    | -77 (-76 to 79)                                   | -51.9    | -140 (-138 to 143)                          | -66.2    | -99 (-97 to 101)                            | -64.2    |
| No known vulnerability | -68 (-67 to 69)                         | -66.5    | -66 (-66 to 67)                                   | -63.2    | -116 (-115 to 118)                          | -75.3    | -84 (-83 to 85)                             | -76.7    |
| Any vulnerability      | -176* (-171 to 182)                     | -51.6    | -124* (-120 to 128)                               | -36.9    | -245* (-237 to 253)                         | -53.2    | -161* (-154 to 168)                         | -47.1    |
| LBW/preterm only       | -69 (-66 to 72)                         | -68.0    | -64 (-62 to 66)                                   | -64.1    | -125* (-121 to 130)                         | -78.3    | -90* (-86 to 94)                            | -79.9    |
| CC only                | -263* (-257 to 270)                     | -47.2    | -157* (-152 to 162)                               | -28.5    | -309* (-300 to 318)                         | -43.1    | -189* (-181 to 197)                         | -34.8    |
| Both                   | -241* (-231 to 251)                     | -59.1    | -216* (-208 to 225)                               | -54.8    | -434* (-419 to 450)                         | -73.1    | -311* (-298 to 325)                         | -72.8    |
| Any CC                 | -259* (-254 to 265)                     | -49.3    | -171* (-166 to 175)                               | -33.0    | -337* (-330 to 345)                         | -48.8    | -217* (-210 to 224)                         | -41.8    |

\*Significantly different from children with no known vulnerability (5% level of significance).  
CC, chronic conditions; LBW, low birth weight.



**Figure 2** Weekly rates of hospital contacts among children aged 0–4 years during the pandemic (March 2020–2021) and averaged for 2015–2019 by age and presence of a chronic condition. Note: lockdown 1: 23 March–23 June 2020; lockdown 2: 5 November–6 December 2020; lockdown 3: 1 January–8 March 2021.

Our study did not examine COVID-19 related contacts because hospitalisation is rare as children typically experience mild asymptomatic disease.<sup>44 45</sup> However, we identified small differences in reductions of hospital care for children in the Asian ethnic group and for children in the most deprived quintile. This suggests that inequalities exacerbated by COVID-19 in adult populations might also extend to children from more deprived, minority ethnic backgrounds.

Potential mechanisms underpinning reductions in planned care likely represent restrictions to access, supported by a rise in postponed outpatient care. Our findings show these restrictions were mitigated in infants, who have a high frequency of hospital care and for whom interventions are likely to be more time critical than in older children.<sup>3</sup> However, reductions remained large in older children and may reflect unmet need or postponed care that could have longer term health consequences.<sup>11</sup> A move to virtual appointments may have exacerbated existing inequalities for families without access to the internet at home.

Reductions in unplanned care may be driven by opposing factors. Previous studies reported decreases in unplanned infection-related hospitalisations due to reduced social exposure and increased hygiene, with little change in admissions for non-infectious causes like appendicitis.<sup>27 30 31 34 41</sup> Others have reported reductions in injury.<sup>46</sup> The spike in unplanned admissions after schools reopened in autumn 2020 and in March 2021 when the third lockdown ended likely reflects increased socialisation. Other positive effects could include reduced exposure to triggers for respiratory disease (eg, air pollution)<sup>47</sup> and improved medication adherence through increased parental supervision. Negative implications could include reduced extrinsic interventions through education, health and social care professionals,<sup>7 48</sup>

or delaying or avoiding medical care due to fears of hospital-acquired COVID-19 infection.<sup>38 49–51</sup> Additionally, these reductions could represent missed opportunities for earlier and more effective intervention.<sup>11 52</sup>

This analysis was the first step in quantifying deferred or foregone hospital care during the pandemic. Studies using routine administrative data report only acute hospital presentations,<sup>37 38</sup> which may reflect late or missed diagnoses.<sup>11 52</sup> Our findings confirm and quantify the reduction in hospital contacts for preschool children in England during the first year of the COVID-19 pandemic. Research is needed to understand reductions in planned care, the types of care, procedures or treatments affected and the short-term and long-term implications for children with specific conditions. More research will be needed to ascertain whether contact rates have returned to previous levels since the end of restrictions. Further research will also be needed to identify vulnerable groups likely to experience adverse outcomes from unmet healthcare in order to target ‘catch-up’ funding and resources to prevent or mitigate these adverse outcomes.<sup>53</sup>

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**Data availability statement** Data may be obtained from a third party and are not publicly available. The data used in this analysis are expected to be available to accredited researchers in 2022 (as part of the ECHILD Database) by applying to the data providers (DfE and NHS Digital).

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#### REFERENCES

- Saxena S, Bottle A, Gilbert R, *et al*. Increasing short-stay unplanned hospital admissions among children in England; time trends analysis '97-06. *PLoS One* 2009;4:e7484.
- Wolfe I, Thompson M, Gill P, *et al*. Health services for children in Western Europe. *The Lancet* 2013;381:1224-34.
- Ruzangi J, Blair M, Cecil E, *et al*. Trends in healthcare use in children aged less than 15 years: a population-based cohort study in England from 2007 to 2017. *BMJ Open* 2020;10:e033761.
- Wijlaars LP, Hardelid P, Woodman J, *et al*. Contribution of recurrent admissions in children and young people to emergency hospital admissions: retrospective cohort analysis of hospital episode statistics. *Arch Dis Child* 2015;100:845-9.
- Harron K, Gilbert R, Cromwell D, *et al*. International comparison of emergency Hospital use for infants: data linkage cohort study in Canada and England. *BMJ Qual Saf* 2018;27:31-9.
- World Health Organisation.. COVID-19 significantly impacts health services for noncommunicable diseases [Internet], 2021. Available: <https://www.who.int/news/item/01-06-2020-covid-19-significantly-impacts-health-services-for-noncommunicable-diseases>
- Isba R, Edge R, Jenner R, *et al*. Where have all the children gone? decreases in paediatric emergency department attendances at the start of the COVID-19 pandemic of 2020. *Arch Dis Child* 2020;105:704.1-704.
- Public Health England. National COVID-19 surveillance reports [Internet], 2021. Available: <https://www.gov.uk/government/publications/national-covid-19-surveillance-reports>
- et alBottle A, Aylin P, Warner M. What happened to English NHS hospital activity during the COVID-19 pandemic? [Internet], 2021. Available: <https://ifs.org.uk/publications/15432>
- Cecil E, Bottle A, Sharland M, *et al*. Impact of UK primary care policy reforms on short-stay unplanned hospital admissions for children with primary Care-Sensitive conditions. *Ann Fam Med* 2015;13:214-20.
- Lynn RM, Avis JL, Lenton S, *et al*. Delayed access to care and late presentations in children during the COVID-19 pandemic: a snapshot survey of 4075 paediatricians in the UK and Ireland. *Arch Dis Child* 2021;106:e8.
- Harron K, Gilbert R, Cromwell D, *et al*. Linking data for mothers and babies in De-Identified electronic health data. *PLoS One* 2016;11:e0164667.
- NHS Digital. Hospital Episode Statistics data changes in 2021 [Internet], 2021. Available: <https://digital.nhs.uk/data-and-information/data-tools-and-services/data-services/hospital-episode-statistics/hospital-episode-statistics-data-changes-in-2021>
- Herbert A, Wijlaars L, Zylbersztejn A, *et al*. Data resource profile: Hospital episode statistics admitted patient care (Hes APC). *Int J Epidemiol* 2017;46:1093-1093i.
- ADMISSION METHOD [Internet], 2022. Available: [https://www.datadictionary.nhs.uk/attributes/admission\\_method.html](https://www.datadictionary.nhs.uk/attributes/admission_method.html)
- Hardelid P, Dattani N, Gilbert R, *et al*. Estimating the prevalence of chronic conditions in children who die in England, Scotland and Wales: a data linkage cohort study. *BMJ Open* 2014;4:e005331.
- Hardelid P, Dattani N, Davey J. Overview of child deaths in the four UK countries [Internet], 2013. Available: <https://www.hqip.org.uk/resource/overview-of-child-deaths-in-the-four-uk-countries/>
- Feudtner C, Christakis DA, Connell FA. Pediatric deaths attributable to complex chronic conditions: a population-based study of Washington state, 1980-1997. *Pediatrics* 2000;106:205-9.
- NHS Digital. Hospital Outpatient Activity [Internet]. Available: <https://digital.nhs.uk/data-and-information/publications/statistical/hospital-outpatient-activity>
- Office of the Deputy Prime Minister, Neighbourhood Renewal Unit. The English indices of deprivation 2004. [Internet]. London: Office of the Deputy Prime Minister, 2004. Available: <https://www.brent.gov.uk/media/323946/SD%2062b%20Indices%20of%20multiple%20deprivation%202004.pdf>
- Ward J, Hargreaves D, Rogers M. Recent and forecast post-COVID trends in hospital activity in England amongst 0 to 24 year olds: analyses using routine hospital administrative data [Internet]. 2021 Feb [cited 2021 Nov 30] p. 2021. Available: <https://www.medrxiv.org/content/10.1101/2021.02.11.21251584v1>
- GOV.UK. Prime Minister's statement on coronavirus (COVID-19): 23 March 2020 [Internet]. Available: <https://www.gov.uk/government/speeches/pm-address-to-the-nation-on-coronavirus-23-march-2020>
- StataCorp. *Stata statistical software*. College Station, TX: StataCorp LLC, 2017.
- Rose KE, Bressan S, *et al*, Group ESGES. Understanding responses of paediatric emergency departments to the first wave of the Covid-19 pandemic – a pan-European perspective. *Arch Dis Child* 2021;106:A161-2.
- Nijman R, Honeyford K, Bressan S. Epidemiology, severity and outcomes of children presenting to emergency departments across Europe during the SARS-COV-2 pandemic: an observational cohort study. *Arch Dis Child* 2021;106:A469-70.
- Scaramuzza A, Tagliaferri F, Bonetti L, *et al*. Changing admission patterns in paediatric emergency departments during the COVID-19 pandemic. *Arch Dis Child* 2020;105:704.2-6.
- Clavenna A, Nardelli S, Sala D, *et al*. Impact of COVID-19 on the pattern of access to a paediatric emergency department in the Lombardy region, Italy. *Pediatr Emerg Care* 2020;36:e597-8.
- Matera L, Nenna R, Ardenti Morini F, *et al*. Effects of relaxed Lockdown on paediatric ER visits during SARS-CoV-2 pandemic in Italy. *Int J Environ Res Public Health* 2021;18. doi:10.3390/ijerph18189547. [Epub ahead of print: 10 Sep 2021].
- Chavasse R, Almaria A, Christopher A, *et al*. The indirect impact of COVID-19 on children with asthma. *Arch Bronconeumol* 2020;56:768-9.
- Angoulvant F, Ouldali N, Yang DD, *et al*. Coronavirus disease 2019 pandemic: impact caused by school closure and national Lockdown on pediatric visits and admissions for viral and nonviral Infections-a time series analysis. *Clin Infect Dis* 2021;72:319-22.
- Kruizinga MD, Peeters D, van Veen M, *et al*. The impact of lockdown on pediatric ED visits and hospital admissions during the COVID-19 pandemic: a multicenter analysis and review of the literature. *Eur J Pediatr* 2021;180:2271-9.
- Williams TC, MacRae C, Swann OV, *et al*. Indirect effects of the COVID-19 pandemic on paediatric healthcare use and severe disease: a retrospective national cohort study. *Arch Dis Child* 2021;106:911-7.

- 33 Dann L, Fitzsimons J, Gorman KM, *et al.* Disappearing act: COVID-19 and paediatric emergency department attendances. *Arch Dis Child* 2020;105:810–1.
- 34 Kishimoto K, Bun S, Shin J-H, *et al.* Early impact of school closure and social distancing for COVID-19 on the number of inpatients with childhood non-COVID-19 acute infections in Japan. *Eur J Pediatr* 2021;180:2871–8.
- 35 Hartnett KP, Kite-Powell A, DeVies J, *et al.* Impact of the COVID-19 Pandemic on Emergency Department Visits - United States, January 1, 2019-May 30, 2020. *MMWR Morb Mortal Wkly Rep* 2020;69:699–704.
- 36 Nolen LD, Seeman S, Bruden D, *et al.* Impact of social distancing and travel restrictions on Non-Coronavirus disease 2019 (Non-COVID-19) respiratory hospital admissions in young children in rural Alaska. *Clin Infect Dis* 2021;72:2196–8.
- 37 Isba R, Edge R, Auerbach M, *et al.* COVID-19: transatlantic declines in pediatric emergency admissions. *Pediatr Emerg Care* 2020;36:551–3.
- 38 Goldman RD, Grafstein E, Barclay N, *et al.* Paediatric patients seen in 18 emergency departments during the COVID-19 pandemic. *Emerg Med J* 2020;37:emermed-2020-210273–7.
- 39 Ferrero F, Ossorio MF, Torres FA, *et al.* Impact of the COVID-19 pandemic in the paediatric emergency department attendances in Argentina. *Arch Dis Child* 2021;106:e5.
- 40 Wirrell EC, Grinspan ZM, Knupp KG, *et al.* Care delivery for children with epilepsy during the COVID-19 pandemic: an international survey of clinicians. *J Child Neurol* 2020;35:924–33.
- 41 Todd IMF, Miller JE, Rowe SL, *et al.* Changes in infection-related hospitalizations in children following pandemic restrictions: an interrupted time-series analysis of total population data. *Int J Epidemiol* 2021;50:1435–43.
- 42 Mathur R, Rentsch CT, Morton CE, *et al.* Ethnic disparities in COVID-19: increased risk of infection or severe disease? – Authors' reply. *The Lancet* 2021;398:390.
- 43 Williamson EJ, Walker AJ, Bhaskaran K, *et al.* Factors associated with COVID-19-related death using OpenSAFELY. *Nature* 2020;584:430–6.
- 44 Ladhani SN, Amin-Chowdhury Z, Davies HG, *et al.* COVID-19 in children: analysis of the first pandemic peak in England. *Arch Dis Child* 2020;105:1180–5.
- 45 Mehta NS, Mytton OT, Mullins EWS, *et al.* SARS-CoV-2 (COVID-19): what do we know about children? A systematic review. *Clinical Infectious Diseases* 2020;71:2469–79.
- 46 Nuñez JH, Sallent A, Lakhani K, *et al.* Impact of the COVID-19 pandemic on an emergency Traumatology service: experience at a tertiary trauma centre in Spain. *Injury* 2020;51:1414–8.
- 47 Krivec U, Kofol Seliger A, Tursic J. COVID-19 lockdown dropped the rate of paediatric asthma admissions. *Arch Dis Child* 2020;105:809–10.
- 48 Kubicek K, Liu D, Beaudin C, *et al.* A profile of nonurgent emergency department use in an urban pediatric hospital. *Pediatr Emerg Care* 2012;28:977–84.
- 49 Czeisler Mark É., Marynak K, Clarke KEN, *et al.* Delay or avoidance of medical care because of COVID-19–Related concerns — United States, June 2020. *MMWR Morb Mortal Wkly Rep* 2020;69:1250–7.
- 50 McDonnell T, Nicholson E, Conlon C, *et al.* Assessing the impact of COVID-19 public health stages on paediatric emergency attendance. *Int J Environ Res Public Health* 2020;17:6719.
- 51 Health Foundation. Public perceptions of health and social care in light of COVID-19, 2020. Available: <https://www.health.org.uk/publications/reports/public-perceptions-of-health-and-social-care-in-light-of-covid-19-may-2020>
- 52 Dayal D, Gupta S, Raihatha D, *et al.* Missing during COVID-19 lockdown: children with onset of type 1 diabetes. *Acta Paediatr* 2020;109:2144–6.
- 53 NHS England » Core20PLUS5 – An approach to reducing health inequalities [Internet], 2021. Available: <https://www.england.nhs.uk/about/equality/equality-hub/core20plus5/>