Prevalence of autistic spectrum disorders in Lothian, Scotland: an estimate using the “capture–recapture” technique

M J Harrison, A E O’Hare, H Campbell, A Adamson, J McNeillage

Aims: To determine the prevalence of autistic spectrum disorder in a geographically defined population for children receiving services and compare this to the estimated prevalence based on a two source “capture-recapture” technique previously employed in biological populations to give a “true” prevalence with full ascertainment.

Methods: Information on gender, age, and postcode sector was determined from nine different datasets maintained for children with autistic spectrum disorder and point prevalence was calculated. Data from the diagnostic services and the Special Needs System were entered into the two source “capture-recapture” calculation.

Results: Of a total population of 134,661 under 15 years old resident in Lothian in southeast Scotland, 443 were known to autism services, with a point prevalence of 32.9 per 10,000 (95% CI 29.8 to 36.0). The estimated prevalence using a capture-recapture method was 44.2 (95% CI 39.5 to 48.9), which suggests that 74% of affected children were registered with services in some way. The age distribution was similar to that of the background population under the age of 12 years and there was no indication of a rising prevalence. The ratio of boys to girls was 7:1.

Conclusions: The prevalence of autistic spectrum disorder in a geographically based population employing two source capture-recapture analysis is comparable to that quoted for the best active ascertainment studies. This technique offers a tool for establishing the prevalence of this condition in health service populations to assist in planning clinical services.
confidentiality and security measures. Nine datasets containing information on ASD cases were identified and questionnaires sent to dataset holders ascertaining catchment area, computerisation, data retention/removal, referral process, inclusion criteria, and definitions of ASD conditions used. Children with autism, ASD, atypical autism, Asperger’s syndrome, and pervasive developmental disorders were included. Only Child and Family Mental Health services who had lost funding for dataset maintenance did not provide information. Other datasets would probably capture these children; as copy correspondence on such children would go to the Community Child Health Department, this could not be confirmed. Only four datasets were computerised.

The dataset covered the previous 10–15 years’ diagnostic clinics throughout Lothian. Diagnoses were based on ICD-10 or DSM-IV pervasive developmental disorders, made using observational assessment by the senior paediatrician or child psychiatrist, with evaluation of communication, reciprocal social interaction, and repetitive behaviours. Attention was paid to the child’s ability to attend, imitate, comprehend, and use language, play appropriately with toys, and interact socially. A range of assessment tools including the Childhood Autism Rating Scale (CARS),18 19 Gilliam Autism Rating Scale,20 and Autism Diagnostic Observation Schedule (ADOS)21 were also employed. Training and practice of senior clinicians using these formal assessments differed throughout Lothian. A E O’Hare reviewed medical notes for all cases to ensure that a senior paediatrician or psychiatrist had made the appropriate diagnosis.

The diagnostic clinic dataset captured children reviewed in child development clinics at secondary and tertiary levels specialising in communication disorder assessment. Diagnosis were usually based on ICD-10 criteria (the coding system for inpatients), but sometimes by DSM-IV (the clinician’s preference). Most children with an ASD received their diagnosis in these clinics, although occasionally in Neurology or Child and Family Mental Health Service Departments. Children seen in the latter were often cross-referred to the Child Development Communication Clinic, especially those of preschool age.

Four further datasets were Lothian-wide. The largest, the Special Needs System, holds information on children with complex disability defined as needing access to two specialised or second tier services within a range of providers including health, education, social, or voluntary services. Diagnosis on the Special Needs System was based on ICD-10 criteria. The pragmatic category of ASD, broadly coinciding with pervasive developmental disorder and including autism and Asperger’s syndrome, described children on these datasets. Essentially, a diagnosis is clinical judgement guided by behavioural symptoms, and ASD is increasingly accepted as encompassing children having significant impairments in communication, social understanding, skill/flexibility in thinking, and behaviour, giving rise to special needs.22 Families can opt out of this system. The remaining datasets contributed to prevalence calculations but were either not Lothian-wide, or targeted narrower age groups or particular clinical needs.

The Lothian Autistic Society, a voluntary organisation, held the final Lothian-wide dataset, and distributed a letter about the study inviting members to provide information if they wanted their child to be included.

Data analysis and statistical tests were conducted using Stata, version 8.2.23 All children from these datasets were included in the prevalence calculation. Point prevalence was used as the estimate was calculated on a specified date, and ASD, despite being of ill defined diagnosis and onset, is stable and lifelong. The total population of under 15 year olds resident in Lothian in mid-2001 was 134 661, based on 2001 population estimates by single age.24

Prevalence calculation for a point estimate:24

- $P = \text{number of existing cases of a condition/total population}$

The likely true number of Lothian children with an ASD was estimated using the two source capture-recapture technique on the best two sources. In this adaptation of the biological sampling capture-recapture technique, databases replace captures, and cases common to both sources represent “tagged” specimens (those captured twice). Four assumptions underpin the two source capture-recapture technique.15 25

Firstly, two comprehensive datasets in a “closed” population (constant population during the study period) must be used. The chance of being referred onto either database must be equal, and datasets must be relatively independent (databases do not refer cases to each other). Finally, cases must be matched confidently and accurately between sources. In practice the two datasets must include all ages of children under 15 years old and cover all of Lothian. Both the diagnostic dataset and Special Needs System satisfied this.

In a two source capture-recapture technique on databases, the “true” number of cases ($N$) is calculated as:12 15 25

$N = [(X+1)(Y+1)/(Z+1)] - 1$

where:

- $X = \text{number of cases from source 1}$
- $Y = \text{number of cases from source 2}$
- $Z = \text{number of cases common to both sources}$

A confidence interval is calculated as:12 25 26

$N \pm 1.96 \sqrt{\text{Var}(N)}$

where $\text{Var}(N) = \left\{ [(X+1)(Y+1)](X)(Y)/(Z+1)^2(Z+2) \right\}$.

The estimated total number of cases can be used in an alternative prevalence estimate of complete ascertainment using the same denominator.25

The apparent trend of smaller proportions of children with ASD in older age groups was tested formally using an extension of the Wilcoxon rank sum test for trend. Children aged less than 3 years ($n = 7$) were excluded; this age group are least likely to be diagnosed with ASD as the condition is more commonly detected once children enter formal education. State-funded nursery is available for those aged over 3 years.

RESULTS

A total of 443 children were identified with ASD (mean age 8.4 years, range 2.5–15 years). Age distributions of children...
with an ASD and the background Lothian population are shown in fig 1. There appeared proportionately reduced numbers of children with ASD in the youngest and oldest ages compared to those aged between 4 and 10 years. There was a significant trend of decreasing proportions of children diagnosed with ASD with increasing age (p = 0.003).

There were 369 boys and 53 girls (11 no data). Table 1 details median age and male:female ratios in each dataset. The median age for Spectrum and SureStart data reflect that these are “early intervention” services for preschool children. The male:female ratio for the Special Needs System was 7:1, but there was an increased ratio of 12.1:1 for the Learning Disability Community Nursing Team service.

Using 443 cases and the denominator of 134 661,\(^2\) prevalence of ASD in Lothian is 32.9 per 10 000 children (95% CI 29.8 to 36.0).

A total of 243 and 268 children with ASD were recorded on the Special Needs System and diagnosis datasets respectively; 109 of these were common to both. An estimated total of 596 children with ASD in Lothian was calculated, using the two source capture–recapture calculation on these two sources.

- \(N = \frac{(243+1)(268+1)}{(109+1)} - 1 = 595.7\) (95% CI 332.4 to 638.9)

Based on the estimate of 596 children with an ASD using the same denominator yielded an estimated prevalence of 44.2 per 10 000 children (95% CI 39.5 to 48.9) can be calculated. This is arguably the true prevalence if complete ascertainment of cases was achieved.

- 596/134661 = 0.00442 × 10000 = 44.2

**DISCUSSION**

We established a point prevalence of 32.9 per 10 000 for children with ASD in Lothian, Scotland, comparable to those quoted by Honda and colleagues and Powell and colleagues.\(^3\) Although the Honda *et al* study used direct ascertainment, it was based on a small population, had wide confidence intervals, and studied childhood autism without widening the concept to ASD. The Powell *et al* study used a larger population and similar approach to this study. We therefore consider that our prevalence figures are likely to be the minimum for childhood ASD and plans for services are likely to be inadequate if provision is for less than this.

However, this study aimed to estimate what “true” prevalence might be, assuming the possibility that not all children were captured on the nine datasets. By applying the “capture–recapture” technique for the first time in ASD, the estimated point prevalence was considerably higher at 44.2 per 10 000. The resulting confidence intervals overlap with the active ascertainment study by Baird and colleagues\(^4\) and approach the lower limits of the Chakrabarti and Fombonne estimate.\(^5\) The former reported on the prevalence of children with ASD, and the latter pervasive developmental disorder (PDD), although it is likely that similar children were included. In Lothian the term ASD has largely replaced PDD and encompasses autism, Asperger’s syndrome and atypical autism. It could be argued that ASD is too inclusive a concept, leading to inflated prevalence figures. However, children in this study were known to services and therefore likely to have significant special needs. It is probably more realistic to plan services for this broader group of children whose central difficulties are in social communication and cognition. Active ascertainment studies such as that in a Welsh education authority have shown a minimum prevalence of 20.2 per 10 000 for children with ASD in mainstream schools; many of these children were unknown to specialist services and had unmet needs for further assessment and management.\(^6\)

Comparison of our two prevalence estimates suggests that only 74% of children with ASD were known to services (identified by any dataset). The age distribution might support the contention that this shortfall was primarily for younger and older children as prevalence across the age groups of 4–12 years was steady. Few children under 3 years old received a diagnosis, consistent with published experience that although parents often retrospectively recognise noticing abnormal interaction and communication in their infant, they had either not sought assessment or had been reassured.\(^7\) The relative lack of older children may be an artefact of diagnostic practice or they may have graduated out of paediatric services. The 13–15 year olds would have received their MMR vaccination as data collection was completed by July 2001. The children’s dates of birth were between July 1986 and June 2001; the MMR was introduced in Lothian in 1987, yet they are less numerous than 4–10 year olds. This would be against an MMR trigger but consistent with greater recognition and acceptance of ASD or better recording as recently suggested.\(^8\)

The “capture–recapture” assumptions must be considered possible explanations for differences between reported and estimated prevalence. Firstly, the “closed” population assumes families are not immigrating/emigrating during the study period. However, estimations suggested Lothian’s population increased by 3.7% between 1991 and 1999.\(^9\) Secondly, children could be accurately matched between datasets; this study matched on initials, gender, date of birth, and postcode, which should satisfy this assumption. Thirdly, children must have equal chance of referral onto either dataset. Doctors and health visitors refer onto both, although health visitor “concern” may be channelled through the child’s general practitioner, speech therapist, or paediatrician for diagnostic clinic referral. Finally, the two systems must be independent of each other. Although there is overlap in referrers, the systems do not automatically refer to each other. There is a degree of dependence of the systems. However, most children are referred to the Special Needs

<table>
<thead>
<tr>
<th>Dataset</th>
<th>N</th>
<th>Median age</th>
<th>Range</th>
<th>Min</th>
<th>Max</th>
<th>M:F ratio</th>
</tr>
</thead>
<tbody>
<tr>
<td>Special Needs System</td>
<td>243</td>
<td>8.4</td>
<td>12.4</td>
<td>2.6</td>
<td>15</td>
<td>7:1</td>
</tr>
<tr>
<td>Diagnosis dataset</td>
<td>268</td>
<td>7.3</td>
<td>12.5</td>
<td>2.5</td>
<td>15</td>
<td>6:2</td>
</tr>
<tr>
<td>Learning Disability Community Team</td>
<td>40</td>
<td>9.6</td>
<td>10.5</td>
<td>4.1</td>
<td>14.6</td>
<td>12.3:1</td>
</tr>
<tr>
<td>Lothian Autistic Society</td>
<td>82</td>
<td>8.4</td>
<td>11</td>
<td>2.9</td>
<td>13.9</td>
<td>6:5:1</td>
</tr>
<tr>
<td>Royal Hospital for Sick Children (hospital dataset)</td>
<td>58</td>
<td>8.2</td>
<td>12.2</td>
<td>2.6</td>
<td>14.8</td>
<td>8:7:1</td>
</tr>
<tr>
<td>Speech and Language Therapy</td>
<td>61</td>
<td>9.6</td>
<td>12.2</td>
<td>2.6</td>
<td>14.8</td>
<td>6:6:1</td>
</tr>
<tr>
<td>VITSS (school age children)</td>
<td>79</td>
<td>9.2</td>
<td>9.6</td>
<td>2.6</td>
<td>13.8</td>
<td>N/A</td>
</tr>
<tr>
<td>Spectrum (preschool children)</td>
<td>21</td>
<td>3.9</td>
<td>2.7</td>
<td>2.6</td>
<td>5.3</td>
<td>8:1:1</td>
</tr>
<tr>
<td>SureStart</td>
<td>54</td>
<td>4.7</td>
<td>3.9</td>
<td>2.6</td>
<td>6.5</td>
<td>9:8:1</td>
</tr>
</tbody>
</table>

*All male*
Autistic spectrum disorders in Scotland

What is already known on this topic

- Autistic spectrum disorders (ASD) are no longer considered rare conditions; they may be the most common serious developmental disabilities of childhood. Active case ascertainment methods—that is, screening, should provide the most accurate prevalence figures by identifying all affected children, but have many practical difficulties including a lack of an adequate screening instrument.

- Capture-recapture techniques have been adapted for use in database based epidemiological studies to estimate a “true” population size; these methods have been used successfully in other conditions.

What this study adds

- The point prevalence of ASD in Lothian using passive ascertainment was 32.9 per 10 000. This prevalence estimate should be considered the minimum rate when planning services. Using the capture-recapture to estimate the total number of children with an ASD, the prevalence estimate was 44.2 per 10 000.

- The capture-recapture technique is a useful tool for establishing the true prevalence of ASD in health services populations for service planning.

System by health visitors. At this point the child is not diagnosed, but likely to be encountered by other services. Once the child is reviewed and formally diagnosed with ASD, the Special Needs System is updated. Theoretically a child would be uncommon. Therefore we consider the Special Needs System is updated. Theoretically a child would be uncommon. Therefore we consider the Special Needs System is updated. Once the child is reviewed and formally diagnosed with ASD, the prevalence estimate was 44.2 per 10 000.

In conclusion, the prevalence of autistic spectrum disorder in Lothian using passive ascertainment studies. The technique offers a tool to estimate a “true” population size; these methods have been used successfully in other conditions. The two point “capture-recapture” analysis compares to those of active ascertainment studies. The technique offers a tool to establishing the prevalence of these conditions in health services populations for service planning.

REFERENCES


Authors’ affiliations
M J Harrison, H Campbell, Dept of Public Health Sciences and Child Life and Health, Reproductive and Developmental Sciences, University of Edinburgh, UK
A E O’Hare, A Adamson, J McNeillage, Royal Hospital for Sick Children, Edinburgh, UK

Competing interests: none declared

www.archdischild.com