Iron deficiency is one of the commonest nutrient deficiencies found in early childhood, which, as well as causing anaemia, has widely recognised neurodevelopmental effects. Iron deficiency may be related to the low iron content of early weaning foods and over-reliance on unmodified cows’ milk. Previous studies have found high levels of anaemia (defined as a haemoglobin below 11 g) in toddlers, particularly those of South Asian origin, but rates can vary widely. In two large population surveys, predominantly in Caucasians, the estimated rate of anaemia at 18 months varied from 12% to 20%. This discrepancy may be explained by the different sampling methods used, but it raises the question of whether a low haemoglobin is truly specific for iron deficiency. A low haemoglobin may also be insensitive, since neurodevelopmental effects have been described with iron deficiency even before the development of anaemia.

There are many other potential markers of iron status which may be used in clinical practice. Markers not commonly used now are serum iron, which mainly reflects dietary iron consumption, and the total iron binding capacity, which has been largely superseded by serum ferritin as a measure of iron stores. The mean cell haemoglobin (MCH) and mean cell volume (MCV) are characteristically reduced in iron deficiency and may drop before total haemoglobin.

Zinc protoporphyrin (ZPP), an early product in the synthesis of haem, rises in iron deficiency, but is not commonly used. Survey work has now established the normal range for these values in healthy toddlers but found that their intercorrelation, though statistically significant, was fairly weak. What is not clear is whether an extreme value of any one marker indicates iron deficiency.

In the absence of a gold standard, the strongest indicator of iron deficiency is a therapeutic response to iron treatment. This study thus set out to examine the extent to which currently used markers of iron deficiency predict a therapeutic response to oral iron, with the hypothesis that a combination of markers would be a better predictor than haemoglobin alone.

METHODS

The subjects for this study were identified from within the Millennium Baby Study, a population based, prospective study of feeding and growth in early childhood, via parental questionnaires and a health check at age 13 months. Here weights and lengths of all children were measured and, where consent was obtained, a venous blood sample was taken for haemoglobin, MCV, and MCH using a Coulter STKS blood cell analyser (Coulter Electronics, Hialeah, FL), ferritin using an electrochemical luminescence immunoassay instrument (Roche Diagnostics, Indianapolis, IN), and ZPP using an Aviv ZP hematoflurometer (Aviv Biomedical, Lakewood, NJ). Ethical approval for the study was obtained from Gateshead LREC.

All children found to have an abnormal value for any one of the measures were offered treatment. The treatment thresholds (table 1) were chosen to accord generally with common clinical practice and to ensure that a reasonable proportion of children would fall below each threshold, as indicated by data for children of similar age in the National Diet and Nutrition Survey (NDNS). Treatment consisted of oral iron (Sytiron (sodium feredetate), 1 ml/kg per day, in three doses) for 6 weeks, in two aliquots. Parents were also given free milk tokens for the duration of the trial, to encourage them to re-start formula milk, as well as a diet sheet listing foods rich in iron. The study nurses kept in contact with families by phone and supplied the second bottle of iron when required. Then 3 months after starting...
treatment, a follow up blood test was taken and all the measures were repeated.

Analysis

For each of the five measures the values before and after treatment were converted to standard deviation (SD) scores ([value-mean]/SD) where the mean and SD were those for that measure at the first screen of the whole cohort. Logged values were used for ferritin and ZPP. Change over time was then calculated for each variable by subtracting the earlier from the later score. A therapeutic response score (TRS) was then calculated by summing the changes in score for each measure. The TRS was not normally distributed, so comparisons between subgroups were made using a non-parametric test (Mann-Whitney).

Demographic information was collected at recruitment and used to dichotomise families into affluent (homeowner, car owner, and one or more employed parent) and deprived (one or more of: rented housing, no car, no employed parent). Information about milk feeding and age of weaning was also available from the parental questionnaires. Weights and heights were converted into SD scores compared to the UK 1990 growth reference.89

RESULTS

Of the 1029 subjects in the main cohort study, 846 (82%) attended the health check and the parents of 495 of those babies agreed for blood to be taken. Samples were obtained from the parental questionnaires. Weights and heights were converted into SD scores compared to the UK 1990 growth reference.89

For most measures only 6–9% of the children had any one value below the treatment threshold (table 1). However 108 (23%) children had values of ZPP above 80 μmol/mol haem. All the markers correlated with each other moderately, except ferritin (table 2). At the time of the health check 108 (23%) children were described as mildly unwell, while three children (0.6%) had a fever. These children had significantly higher ferritin levels than the remainder (geometric mean unwell or fever 28.0, remainder 21.0; p = 0.001, x² test). ZPP levels were unaffected.

In total 147 (32%) children had at least one abnormal value and were offered treatment. Milk tokens were issued to 125 families, but they tended to have difficulty persuading their children to re-start formula milk. At follow up all measures tended to have improved (table 3), but there was a wide variation and the median change for any one variable was small. Paired ferritin values were only available for 104 (82%) subjects and ferritin showed the weakest correlation with other values, so the TRS

<table>
<thead>
<tr>
<th>Measure</th>
<th>Number</th>
<th>Median</th>
<th>IQR</th>
<th>Median</th>
<th>IQR</th>
</tr>
</thead>
<tbody>
<tr>
<td>Haemoglobin (g/dl)</td>
<td>126</td>
<td>0.00</td>
<td>-0.5 to 0.71</td>
<td></td>
<td></td>
</tr>
<tr>
<td>MCV (fl)</td>
<td>126</td>
<td>0.40</td>
<td>-0.25 to 0.60</td>
<td></td>
<td></td>
</tr>
<tr>
<td>MCH (pg)</td>
<td>126</td>
<td>0.55</td>
<td>-0.15 to 0.54</td>
<td></td>
<td></td>
</tr>
<tr>
<td>Ferritin (μg/l)</td>
<td>104</td>
<td>0.5</td>
<td>0.01 to 0.96</td>
<td></td>
<td></td>
</tr>
<tr>
<td>ZPP (μmol/mol haem)</td>
<td>125</td>
<td>-15.0</td>
<td>-0.62 to 1.18</td>
<td></td>
<td></td>
</tr>
</tbody>
</table>

*Sum of changes in standardised scores for Hb, MCV, MCH, and ZPP.
was calculated using only the other four markers, to maximise the number of children who could be included in the analysis. The TRS on average also showed a net gain but with a wide range (table 3); 64% of children had a TRS greater than 0 with 49% greater than 1. The TRS was no higher in those receiving 6 as opposed to 3 weeks treatment.

Children with a haemoglobin or MCH below the screening cut off showed the largest therapeutic response to iron (table 4). Children with values above the screening threshold for ZPP (80) did not generally show a therapeutic response to iron. However those above a higher threshold of 97, found in 10% of the original cohort, did show a reasonable response. Some 44 children had either a low haemoglobin or a low MCH or both, and they showed a median TRS of 1.18. A further 11 children had abnormal values for two or more of the remaining three measures. All these children together showed a median TRS of 1.68. The remaining 70 children who had normal haemoglobin, MCH, and only one other marker abnormal, showed a median TRS of only 0.36. In the original screened population as a whole, 13% of children had abnormal haemoglobin or MCH or abnormal levels of two or more of the remaining three markers and could be defined as iron deficient.

Rates of iron deficiency were slightly higher in the more deprived children, infants introduced to cows’ milk before 12 months, and those starting solids or continuing breast feeding after the age of 4 months, and much higher in a small group of infants from an ultra Orthodox Jewish group (3% cohort) with very different feeding patterns from the majority population. However, only the latter two factors were independently statistically significant (table 5). Children with iron deficiency were slightly lighter (mean difference 0.22), but the differences were not significant (p = 0.25 and 0.14, respectively).

**DISCUSSION**

The diagnosis of severe iron deficiency is straightforward when a low haemoglobin is accompanied by microcytosis and hypochromia. The difficulty is interpreting low markers in isolation, particularly if the haemoglobin is normal. Our findings reinforce the importance of a low haemoglobin and, unexpectedly MCH, as specific markers of iron deficiency in this age group. However, these are still insensitive, as alone each only identified half the children who showed a substantial treatment response. The MCV and ZPP were much less specific, but when both were abnormal these also produced a substantial therapeutic response, even if the haemoglobin and MCH were within normal limits. Ferritin correlated only weakly with the other markers probably because of its strong tendency to rise in the presence of any infection. Zinc protoporphyrin is not often used clinically and initially seemed highly non-specific, but it became apparent that we used different assay techniques from the NDNS, on which our original threshold was based. Using the more extreme threshold ZPP proved a fair predictor, as has been suggested previously, and certainly better than ferritin.

The purpose of the study was to study change over time within the cohort, rather than describe the prevalence of iron deficiency, so it was not a major problem that less than half the cohort agreed to take part, due mainly to parental reluctance to distress their child. However, though those who did take part were slightly more affluent, they were still broadly representative of the general population.

The challenge to any study of iron deficiency is the absence of a gold standard test, and a therapeutic response to iron is theoretically the best indicator. Abnormal test results will always tend to improve on re-measurement, but the TRS, by standardising and combining responses, allows comparison both of the size of the response and the extent to which other markers have changed.

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**Table 5** Association between dietary and social factors and iron deficiency

<table>
<thead>
<tr>
<th>Measure</th>
<th>Yes</th>
<th>No</th>
<th>Crude OR</th>
<th>p†</th>
<th>Adjusted OR</th>
<th>p‡</th>
</tr>
</thead>
<tbody>
<tr>
<td>Deprived</td>
<td>15.7% (34)</td>
<td>10.6% (26)</td>
<td>1.56</td>
<td>0.07</td>
<td></td>
<td></td>
</tr>
<tr>
<td>Weaned &gt; 4 months</td>
<td>21.6% (8)</td>
<td>11.4% (32)</td>
<td>2.14</td>
<td>0.074</td>
<td></td>
<td></td>
</tr>
<tr>
<td>Breast fed &gt; 4 months</td>
<td>24.4% (19)</td>
<td>10.7% (41)</td>
<td>2.69</td>
<td>0.002</td>
<td>2.44</td>
<td>0.005</td>
</tr>
<tr>
<td>Cows’ milk ≤ 4 months</td>
<td>14.8% (18)</td>
<td>12.4% (42)</td>
<td>1.23</td>
<td>0.30</td>
<td></td>
<td></td>
</tr>
<tr>
<td>Ultra Orthodox Jewish†</td>
<td>57.4% (4)</td>
<td>12.3% (56)</td>
<td>9.32</td>
<td>0.007</td>
<td>6.99</td>
<td>0.015</td>
</tr>
</tbody>
</table>

*Logistic regression, all variables p < 0.1 on univariate analysis added initially, then only variables p < 0.05 left in model.
†Pearson’s x² test
‡Logistic regression
†A small group of families with weaning and other feeding practices very different from the remainder of the cohort.
Values are proportion (%).
Compliance with iron therapy is a major problem in clinical practice and it is difficult to estimate how good it was in this cohort. Some 40% percent of the children received only 3 weeks of treatment but showed no difference in response to those receiving the full course. An unexpected finding was the high rates of diarrhoea, particularly in those not completing their treatment course, and this possibly relates to the use of sorbitol as a sweetening agent.

This population showed a much lower prevalence of anaemia than a previous study in Bristol, although their mean ferritin levels were higher. The Bristol study used capillary sampling, which recent work suggests may substantially underestimate true haemoglobin levels. We also found slightly lower rates of anaemia than the NDNS, but in our study the children were 5 months younger and the majority had only recently stopped formula milk, which is strongly fortified with iron. Because of the relatively low prevalence, we had limited power to explore possible risk factors. We found little association between iron deficiency and early consumption of cows’ milk, unlike an intervention study where deferring unmodified cows’ milk until after the age of a year had a significant effect in preventing iron deficiency.

We found only a weak non-significant trend towards higher rates of iron deficiency in children from the more deprived families, which is in keeping with the NDNS.

We did find an excess of iron deficiency in the minority of children breast fed for more than 4 months, which mirrors the findings of a recent trial and further endorses the recommendation that breast fed infants need additional sources of iron after the age of 6 months.

CONCLUSIONS

These findings suggest that in toddlers low haemoglobin or low MCH are highly specific for iron deficiency and should always merit treatment. Low MCV, high ZPP, and possibly low ferritin levels if found in combination also merit treatment, even if the haemoglobin and MCH are normal, but not if found in isolation. Diarrhoea is a common side effect of iron treatment that may limit parental compliance.

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